

Efficient Techniques, MC Revisited

Monte Carlo Method

Applicable to any system, Unbiased, Consistent

Long Execution time

Semi Analytic Technique to Accelerate MC + Method of Moments

Considerable acceleration

Needs the pdf of sufficient statistics of the symbol detection decision

Other Techniques to overcome lengthy run-time of MC

Tail Extrapolation

Curve fitting to MC simulation results

PDF Estimators

Model based estimating the pdf of decision metric

Importance Sampling

Biasing the channel noise in a controlled way to increase error occurrence

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Tail Extrapolation

Extrapolation → Making inferences beyond the range of available data

Collecting data in higher error rates and extrapolate to low error rate cases

Higher Error Rates regimes are quick to simulate

Tail distribution assumption is class of Generalized Exponential (GE) distributions

$$f(x|v, \mu, \sigma) = \frac{v}{2\sqrt{2}\sigma \Gamma(1/v)} \exp\left\{-\left|\frac{x-\mu}{\sqrt{2}\sigma}\right|^v\right\}$$

$$\Gamma(a) = \int_0^\infty t^{a-1} e^{-t} dt \rightarrow \Gamma(n) = (n-1)!$$

For error threshold = T:

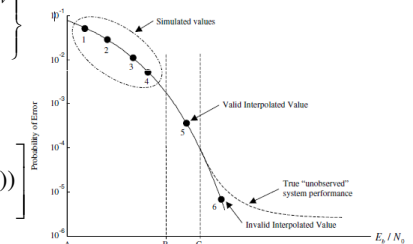
$$\int_{T+\mu}^\infty f(x|v, \mu, \sigma) dx \approx \exp\left[-\left(\frac{T}{\sigma\sqrt{2}}\right)^v (1-\epsilon(T))\right]$$

$$\epsilon(T) \ll 1 \rightarrow 1 - \epsilon(T) \approx 1$$

$$\ln(-\ln p(T)) = v \ln(T/\sqrt{2}) - v \ln(\sigma)$$

For M threshold values T_i , estimate $p(T_i)$ using MC

Using linear regression or LMS we can fit a line to the data pairs $(T_i, p(T_i))$



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Tail Extrapolation...

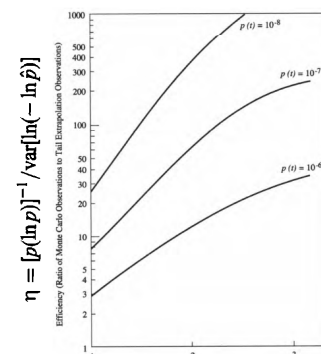
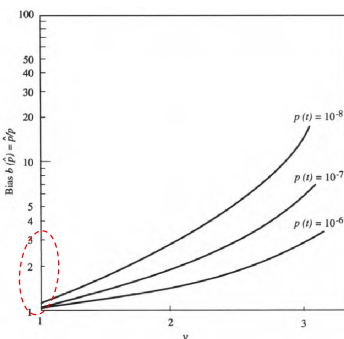
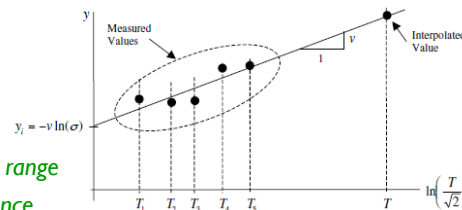
From slope → v

From intersect and v → σ

Estimation is biased

Bias can be kept in a reasonable range

Trade-off between bias and variance



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PDF Estimators

Having the pdf of decision statistics → $P_E = \int_T^\infty f_V(v) dv$

Histogram method is the most straight forward → biased, needs data for tails

Parametric methods are more robust

Gram-Charlier Series

a posteriori method of approximating an unknown, but nearly Gaussian, pdf with known moments: $E[Y^k]$

$$\hat{f}_Y(y) = \frac{1}{\sigma_Y \sqrt{2\pi}} \exp\left[-\frac{(y-\mu_Y)^2}{2\sigma_Y^2}\right] \sum_{j=1}^{\infty} C_j H_j\left(\frac{y-\mu_Y}{\sigma_Y}\right), \quad C_k = \frac{1}{k!} \int_{-\infty}^{\infty} H_k(y) f_Y(y) dy$$

$$H_k(y) = \frac{[-\exp(-y^2/2)]^{(k)}}{\exp(-y^2/2)} \quad C_k = \frac{1}{k!} \sum_{i=0}^{k/2} \frac{k^{2i}}{2^i i!} (-1)^i \mu_{k-2i} \quad \mu_k = E\left[\frac{Y-\mu_Y}{\sigma_Y}\right]^k$$

$$H_k(y) = yH_{k-1}(y) - (k-1)H_{k-2}(y), \quad k \geq 2 \quad H_0(y) = 1, \quad H_1(y) = y$$

H: Hermite Polynomials If only two moments are known → Gaussian pdf

Good approximation around mean → biased, no reasonable truncation

Yields excellent results in some wireless channel conditions

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PDF Estimators...

Parzen Series

provides a smoothed estimator of a pdf

Part of a bigger class of kernel estimators

$$\hat{f}_Y(y) = \hat{f}_{Y|Y_1, \dots, Y_n}(y | y_1, \dots, y_n) = \frac{1}{nh(n)} \sum_{i=1}^n g\left(\frac{y - y_i}{h(n)}\right)$$

h : smoothing function, g : weighting function, n : data size

Asymptotically unbiased and consistent

For Gaussian like pdfs reasonable choices for g and h are:

$$h(n) = 1/\sqrt{n}, \quad g(x) = \frac{1}{\sqrt{2\pi}} \exp(-x^2/2)$$

$$\rightarrow \hat{f}_Y(y) = \frac{1}{\sqrt{2\pi n}} \sum_{i=1}^n \exp(-n(y - y_i)^2/2)$$

Number of terms = number of samples

Has shown good results for multipath + AWGN problems

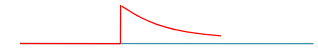
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Importance Sampling

Rare events are important! \leftarrow in MC

For low probabilities, a rare event produces a peak with $1/N$ vanishing rate

$$\lim_{n \rightarrow \infty} \frac{N_A}{N} = P(A)$$



Evaluation of expected values needs rare events or extreme events to happen

For instance for BER we can bias the noise to make rare events happen more frequently

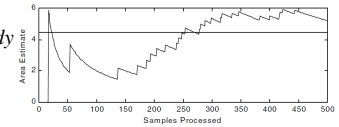
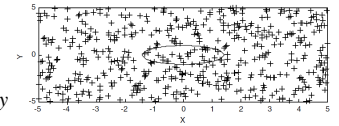
Example: Ellipse Area Estimation

$$x^2 + 2y^2 < 2 \quad A = \int_{-1}^1 \int_{-\sqrt{2-2y^2}}^{\sqrt{2-2y^2}} dx dy = \pi\sqrt{2} = 4.443$$

$$h(x, y) = 0 \text{ outside the ellipse, } 1 \text{ inside: } A = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(x, y) dx dy$$

$$A = A_b \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(x, y) (1/A_b) dx dy = A_b \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(x, y) f_{XY}(x, y) dx dy$$

$$A = A_b \left(\frac{N_e}{N} \right)$$



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Importance Sampling...

Example: Ellipse Area Estimation...

Performing MC for different bounding sizes

Box size decreases, so does the error, but down to a limit...

Optimal bounding region, for uniform f_{XY}

The ellipse itself! Converges to true value with just the first sample! \rightarrow not a solution!

Smallest possible bounding rectangle or even a polygon \rightarrow may be found analytically

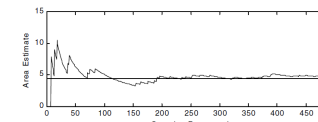
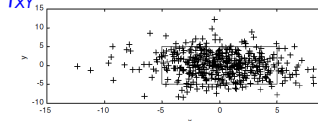
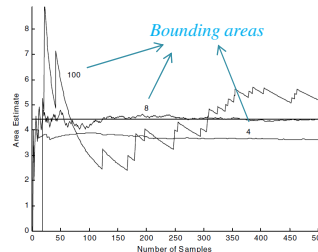
Rectangle is attractive because of the simplicity of f_{XY}

Any (almost) non-uniform pdf can be used,

$$f_{XY}(x, y) = \begin{cases} 1/w(x, y), & h(x, y) = 1 \\ \text{arbitrary}, & h(x, y) = 0 \end{cases}$$

$$A = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(x, y) w(x, y) f_{XY}(x, y) dx dy$$

$$\hat{A} = \frac{1}{N} \sum_{i=1}^N h(x_i, y_i) w(x_i, y_i) \quad (x_i, y_i) \sim f_{XY}(x, y)$$



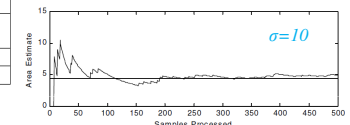
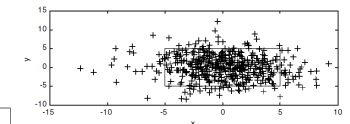
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Importance Sampling...

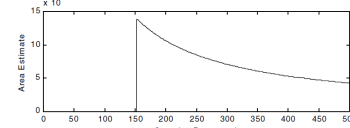
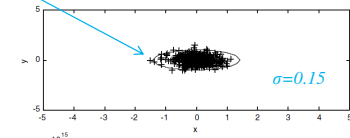
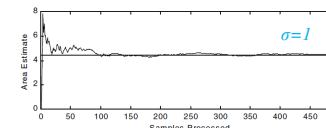
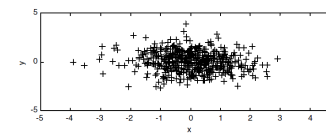
Example: Ellipse Area Estimation...

Gaussian, $\sigma = 10, 1, 0.15$

| Standard Deviation, σ | Number of Samples in Ellipse | Mode of Convergence |
|------------------------------|------------------------------|--------------------------------|
| Too small | Nearly all | Slow (exhibits extreme events) |
| Appropriate | Between 10% and 90% | Rapid |
| Too large | Very few | Slow |



Extreme Event

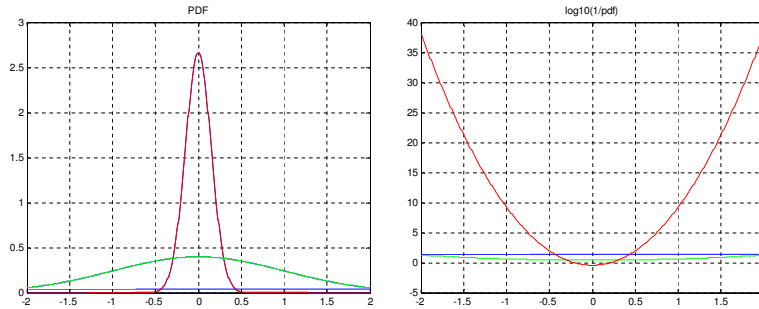


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Importance Sampling...

Gaussian, $\sigma=10, 1, 0.15$

$\sigma=10$ $\sigma=1$ $\sigma=0.15$



| Standard Deviation, σ | Number of Samples in Ellipse | Mode of Convergence |
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| Too small | Nearly all | Slow (exhibits extreme events) |
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Importance Sampling... Sensitivity to pdf choice

If the random data comes from a physical source, like noise in communication systems

$$B = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(x, y) f_{phy}(x, y) dx dy \xrightarrow{MC} \hat{B} = \frac{1}{N} \sum_{i=1}^N h(x_i, y_i), \quad (x_i, y_i) \sim f_{phy}, IID$$

B is like the probability that the generated points fall inside the ellipse

And we use a different noise, called biased noise, say f_{sim}

$$B = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(x, y) w(x, y) f_{sim}(x, y) dx dy \xrightarrow{MC} \hat{B} = \frac{1}{N} \sum_{i=1}^N h(x_i, y_i) w(x_i, y_i)$$

$$(x_i, y_i) \sim f_{sim}, IID \quad w(x, y) = \begin{cases} f_{phy}(x, y) / f_{sim}(x, y), & h(x, y) = 1 \\ \text{arbitrary}, & h(x, y) = 0 \end{cases}$$

Easy to show that sufficient condition for having unbiased and consistent estimation is

$$f_{sim}(x, y) > 0 \text{ for all } (x, y) \text{ that } h(x, y) f_{phy}(x, y) > 0$$

Mode of convergence depends on $w(x, y)$

| PDFs Inside Ellipse | Weighting Function Inside Ellipse | Rate of Convergence |
|---------------------------------|-----------------------------------|-----------------------|
| $f_{sim}(x, y) = f_{phy}(x, y)$ | $w(x, y) = 1$ | Same as simple MC |
| $f_{sim}(x, y) > f_{phy}(x, y)$ | $w(x, y) < 1$ | Faster than simple MC |
| $f_{sim}(x, y) < f_{phy}(x, y)$ | $w(x, y) > 1$ | Slower than simple MC |

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Importance Sampling... Applied to Communication Systems

Channel noise goes through the receiver block and its pdf changes!

Unknown or difficult to track pdf

Assume that channel generates $(x, y) \sim f_{phy}$

Receiver maps (x, y) to $g(x, y) = (a, b) \sim f_{rec} \rightarrow$ impractical to calculate f_{rec} from f_{phy}

$$B = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(a, b) f_{rec}(a, b) da db \xrightarrow{MC} \hat{B} = \frac{1}{N} \sum_{i=1}^N h(a_i, b_i), \quad (a_i, b_i) \sim f_{rec}, IID$$

$$\hat{B} = \frac{1}{N} \sum_{i=1}^N h(g(x_i, y_i)) \quad (x_i, y_i) \sim f_{phy}$$

Now alter f_{phy} to f_{bias} and to remove the effect of this bias from MC estimation use a weighting function...

$$w(a, b) = \frac{f_{rec}(a, b)}{f_{rec-bias}(a, b)} = \frac{f_{phy}(x, y)}{f_{bias}(x, y)} \leftarrow (a, b) = g(x, y)$$

$$\hat{B} = \frac{1}{N} \sum_{i=1}^N w(g(x_i, y_i)) h(x_i, y_i)$$

Not dependent on f_{rec}

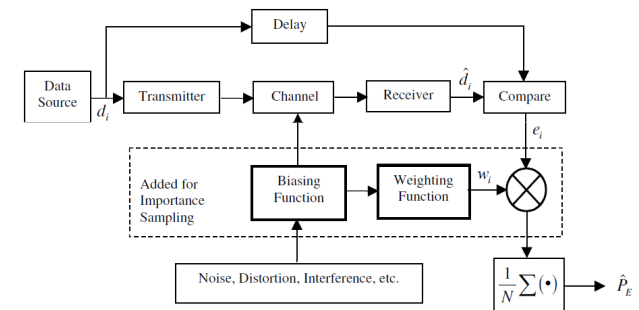
Proved to be unbiased and consistent

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Importance Sampling... Applied to Communication Systems...

Mapping the MC area estimation to BER(SER) estimation

$$w_i = f_{phy}(n_i) / f_{bias}(n_i) \rightarrow \hat{p}_e = \frac{1}{N} \sum_{i=1}^N w_i h(n_i)$$



Ideally desired is to have a mixture of samples inside and outside of error region

Too many samples in either region will cause the convergence rate to be very slow

CIS and IIS \rightarrow conventional and improved importance sampling

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Importance Sampling... Applied to Communication Systems...

CIS: changing the variance (IS10)

$$w_i = f_{phy}(n_i) / f_{bias}(n_i) = \frac{\sigma_b}{\sigma_p} \exp(-n_i^2 (\sigma_p^{-2} - \sigma_n^{-2}) / 2)$$

IIS: Shifting the pdf (IS9)

$$w_i = f_{phy}(n_i) / f_{bias}(n_i) = \exp((2n_i c - c^2) / 2\sigma^2)$$

AIS: Adaptive IS

$$p_e = E(g(x)) \xrightarrow{IS} \hat{p}_e = \frac{1}{N} \sum_{i=1}^N g(x_i) w(x_i, \theta) \quad x_i \sim f_b(x, \theta)$$

$$\text{var}(\hat{p}_e) = \frac{1}{N} (I(\theta) - p_e^2) \leftarrow I(\theta) = E(g^2(x) w(x, \theta)) = E_b(g^2(x) w^2(x, \theta))$$

g is the indicator of $\{x > T\}$ for BER estimation problem

θ can be optimized on each epoch to achieve the best results... (IS11)

Papers that compare several pdf classes for different applications... (IS3)

Optimum and sub optimum parameters (IS4)

EX8:
16.1, 2, 6, 9, 10, 11